# Feature Selection and Sparsity

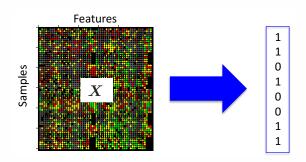
Makoto Yamada
makoto.yamada@oist.jp
Okinawa Institute of Science and Technology

- 1 Introduction
- 2 Feature Selection Algorithms

## Introduction

Feature selection is important for high-dimensional data:

- User data (d > 100), e.g., e-mail spam detection.
- Gene expression data (d > 20000), e.g., cancer classification.
- Text based feature such as TF-IDF (d > 100,000)



#### Motivation1

#### The purpose of feature selection is

- to improve the prediction accuracy by getting rid of non-important features.
- to make the prediction faster.
- to interpret data.
- to handle high-dimensional data.

#### Motivation2

Let us think about the least-squared regression problem:

$$\min_{oldsymbol{w} \in \mathbb{R}^d} \ \|oldsymbol{y} - oldsymbol{X}^ op oldsymbol{w}\|_2^2$$

where 
$$\boldsymbol{x} = (x_1, x_2, \dots, x_d)^{\top} \in \mathbb{R}^d$$
,  $\boldsymbol{X} = (\boldsymbol{x}_1, \boldsymbol{x}_2, \dots, \boldsymbol{x}_n) \in \mathbb{R}^{d \times n}$ ,  $\boldsymbol{w} = (w_1, w_2, \dots, w_d)^{\top} \in \mathbb{R}^d$ ,  $\boldsymbol{y} \in \mathbb{R}^n$ , and  $\|\cdot\|_2^2$  is the  $\ell_2$  norm.

#### Question:

• d < n and the rank of  $\boldsymbol{X}$  is d. Please derive the analytical solution of  $\boldsymbol{w}$ .

## **Motivation2**

Take the derivative with respect to  $\boldsymbol{w}$  and set it to zero:

$$\frac{\partial}{\partial \boldsymbol{w}} \|\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}\|_{2}^{2} = -2\boldsymbol{X}(\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}) = \boldsymbol{0}$$

Use Eq. (84) of [1]. The solution is given as

$$\widehat{\boldsymbol{w}} = (\boldsymbol{X} \boldsymbol{X}^{\top})^{-1} \boldsymbol{X} \boldsymbol{y}.$$

If the rank of X is d,  $XX^{\top}$  is invertible.

What happens if the rank of X is less than d?

•  $XX^{\top}$  is not invertible.

A possible solution is to use feature selection! If we select r < d features, we can compute  ${\boldsymbol w}.$ 

#### **Problem formulation**

Problem formulation of feature selection:

- Input vector:  $\boldsymbol{x} = (x_1, x_2, \dots, x_d)^{\top} \in \mathbb{R}^d$
- Output:  $y \in \mathbb{R}$
- Paired data:  $\{(\boldsymbol{x}_i, y_i)\}_{i=1}^n$

Goal: Select r(r < d) features of input x that are responsible for output y.

Problems: There is  $2^d$  combinations :( It is hard even if d is 100.

- 1 Introduction
- 2 Feature Selection Algorithms

## Feature Selection Algorithms

The feature selection algorithms can be categorized into three types.

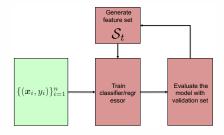
- Wrapper Method
   Use a predictive model to select features.
- Filter Method
   Use a proxy measure (such as mutual information)
   instead of the error rate to select features.
- Embedded Method
   Features are selected as part of the model construction process.

## Wrapper Method

Use a predictive model (e.g., classifier) to select features.

The simplest approach would be...

- **1** Generate feature set  $\mathcal{S}_t$
- 2 Train predictive model with  $S_t$  and test the prediction accuracy with hold-out set.
- 3 Iterate 1 and 2 until all feature combination is examined.



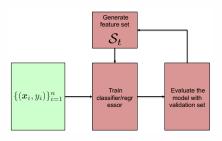
## Wrapper Method

#### Pro:

• It can select features that have feature-feature interaction.

#### Cons:

• Computationally expensive  $(2^d$  combination).



### Filter Method

Use a proxy measure (such as mutual information) instead of the error rate to select features.

#### Pros:

- It scales well.
- Can select features from high-dimensional data (both linear and nonlinear way).

#### Cons:

- The feature selection is independent of the model. The selected features may not be the best set to achieve highest accuracy.
- It is hard to detect select features with interaction.

## Filter Method (Example)

#### Maximum Relevance Feature Selection (MR)

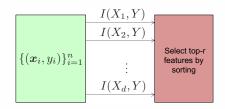
Compute association score between each feature and its output and rank them.

- Correlation, Mutual information, and the kernel based independence measures are used.
- Easy to implement and it scales well.

#### Optimization problem:

$$\max_{\beta \in \{0,1\}^d} \frac{1}{S} \sum_{k=1}^d \beta_k I(X_k, Y),$$

where 
$$S = \beta_1 + \ldots + \beta_d$$
.



## Filter Method (Example)

#### Minimum Redundancy Maximum Relevance (mRMR) [2]

MR feature selection tends to select redundant features.

#### mRMR method is to

- select features that have high association to its output.
- select independent features.

#### Optimization problem:

$$\max_{\beta \in \{0,1\}^d} \ \frac{1}{S} \sum_{k=1}^d \beta_k I(X_k, Y) - \frac{1}{S^2} \sum_{k=1}^d \sum_{k'=1}^d \beta_k \beta_{k'} I(X_k, X_{k'}).$$

This optimization problem can be solved by using greedy algorithm.

## Filter Method (Mutual Information)

To optimize mRMR, we tend to use the mutual information as an association score.

#### Independence:

$$p(\boldsymbol{x}, \boldsymbol{y}) = p(\boldsymbol{x})p(\boldsymbol{y})$$

#### Mutual Information:

$$\mathsf{MI}(X,Y) = \iint p(\boldsymbol{x},\boldsymbol{y}) \log \frac{p(\boldsymbol{x},\boldsymbol{y})}{p(\boldsymbol{x})p(\boldsymbol{y})} \mathsf{d}\boldsymbol{x} \mathsf{d}\boldsymbol{y}$$

Under independence:

$$\mathsf{MI}(X,Y) = \iint p(m{x},m{y}) \log rac{p(m{x})p(m{y})}{p(m{x})p(m{y})} \mathsf{d}m{x} \mathsf{d}m{y} = 0$$

## Filter Method (Linear Correlation)

To optimize mRMR, we may be able to use the Pearson's correlation coefficient

Pearson's correlation coefficient:

$$PCC(X,Y) = \frac{Cov(X,Y)}{\sigma_X \sigma_Y},$$

$$Cov(X,Y) = \mathbb{E}[(X - \mu_X)(Y - \mu_Y)]$$

where  $\mu_X = \mathbb{E}[X]$ ,  $\mu_Y = \mathbb{E}[Y]$ ,  $\sigma_X^2 = \mathbb{E}[(X - \mu_X)^2]$ , and  $\sigma_Y^2 = \mathbb{E}[(Y - \mu_Y)^2]$ .

The cross-covariance can be written as

$$Cov(X,Y) = \mathbb{E}[(X - \mu_X)(Y - \mu_Y)] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y].$$

That is, if PCC(X, Y) = 0,  $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ 

## The relationship between independence and correlation

If X and Y are independent, we can write

$$\mathbb{E}[XY] = \iint xy \ p(x,y) dx dy,$$

$$= \iint xy \ p(x)p(y) dx dy, (independence)$$

$$= \left( \int x \ p(x) dx \right) \left( \int y \ p(y) dy \right)$$

$$= \mathbb{E}[X]\mathbb{E}[Y]$$

That is, if X and Y are independent,  $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ . Note that, even if  $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$ , X and Y can be dependent.

## **Empirical estimation of Cross-covariance**

To optimize mRMR, we may be able to use the Pearson's correlation coefficient

Cross-Covariance (population):

$$Cov(X,Y) = \mathbb{E}[(X - \mu_X)(Y - \mu_Y)]$$

Cross-Covariance estimation:

$$\widehat{\mathsf{Cov}}(X,Y) = \frac{1}{n} \sum_{i=1}^{n} (x_i - \widehat{\mu}_X)(y_i - \widehat{\mu}_Y)$$

$$\widehat{\mu}_X = \frac{1}{n} \sum_{i=1}^{n} x_i = \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \mathbf{1}_n, \quad \widehat{\mu}_Y = \frac{1}{n} \sum_{i=1}^{n} y_i = \frac{1}{n} \boldsymbol{y}^{\mathsf{T}} \mathbf{1}_n,$$

where  $\mathbf{1}_n = (1, 1, \dots, 1)^{\top} \in \mathbb{R}^n$  is the vector with all ones.

## **Empirical estimation of cross-covariance**

#### Cross-Covariance estimation:

$$\widehat{\mathsf{Cov}}(X,Y) = \frac{1}{n} \sum_{i=1}^{n} (x_i - \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \mathbf{1}_n) (y_i - \frac{1}{n} \boldsymbol{y}^{\mathsf{T}} \mathbf{1}_n)$$

$$= \frac{1}{n} \left( \sum_{i=1}^{n} x_i y_i - \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \mathbf{1}_n \mathbf{1}_n^{\mathsf{T}} \boldsymbol{y} \right)$$

$$= \frac{1}{n} \left( \boldsymbol{x}^{\mathsf{T}} \boldsymbol{y} - \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \mathbf{1}_n \mathbf{1}_n^{\mathsf{T}} \boldsymbol{y} \right)$$

$$= \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \left( \boldsymbol{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^{\mathsf{T}} \right) \boldsymbol{y}$$

$$= \frac{1}{n} \boldsymbol{x}^{\mathsf{T}} \boldsymbol{H} \boldsymbol{y},$$

where  $H = I_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^{\top}$  is the centering matrix and  $I_n$  is the identity matrix. (Note HH = H).

## **Empirical estimation of covariance**

#### Covariance estimation:

$$egin{aligned} \widehat{\mathsf{Cov}}(X,Y)^2 &= rac{1}{n^2} oldsymbol{x}^ op oldsymbol{H} oldsymbol{y} oldsymbol{x}^ op oldsymbol{H} oldsymbol{y} oldsymbol{x}^ op oldsymbol{H} oldsymbol{y} oldsymbol{T} oldsymbol{H} oldsymbol{y}^ op oldsymbol{Y} oldsymbol{H} oldsymbol{y}^ op oldsymbol{H} oldsymbol{Y} oldsymbol{H} oldsymbol{y}^ op oldsymbol{H} oldsymbol{Y} oldsymbol{Y} oldsymbol{H} oldsymbol{Y} oldsymbol{H} oldsymbol{Y} oldsymbol{Y} oldsymbol{Y} oldsymbol{Y} oldsymbol{H} oldsymbol{Y} oldsymbol{H} oldsymbol{Y} old$$

where  $m{K} = m{x}m{x}^{ op} \in \mathbb{R}^{n imes n}$  and  $m{L} = m{y}m{y}^{ op} \in \mathbb{R}^{n imes n}.$ 

## Advanced Topic (Hilbert-Schmidt Independence Criterion)

Hilbert Schmidt Independence Criterion (HSIC) [3]

Empirical V-statistics of HSIC is given as

$$\mathsf{HSIC}(X,Y) = \frac{1}{n^2} \mathsf{tr}(\boldsymbol{KHLH}),$$

where we use the Gaussian kernel:

$$oldsymbol{K}_{ij} = \exp\left(-rac{\|oldsymbol{x}_i - oldsymbol{x}_j\|_2^2}{2\sigma^2}
ight), \quad oldsymbol{L}_{ij} = \exp\left(-rac{\|oldsymbol{y}_i - oldsymbol{y}_j\|_2^2}{2\sigma^2}
ight).$$

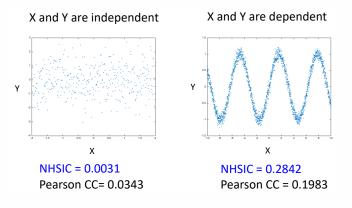
HSIC takes 0 if and only if X and Y are independent.

Since we can decompose  $K = \Phi^ op \Phi$  and  $L = \Psi^ op \Psi$ , we have

$$\mathsf{HSIC}(X,Y) = \frac{1}{n^2}\mathsf{tr}(\boldsymbol{\Phi}^{\top}\boldsymbol{\Phi}\boldsymbol{H}\boldsymbol{\Psi}^{\top}\boldsymbol{\Psi}\boldsymbol{H}) = \frac{1}{n^2}\|\mathsf{vec}(\boldsymbol{\Psi}\boldsymbol{H}\boldsymbol{\Phi}^{\top})\|_2^2 \geq 0$$

## Advanced Topic (HSIC)

Hilbert-Schmidt Independence Criterion (HSIC) experiments



### **Embedded Method**

Features are selected as part of the model construction process. Embedded method can be regarded as an intermediate method between wrapper and filter methods.

#### Pros:

- Can select features with high prediction accuracy.
- Computationally efficient than wrapper method.

#### Cons:

- Computationally expensive than filter method.
- If the input output relationship are nonlinear, it is computationally expensive. It is more suited for linear method.

## **Embedded Method (Lasso)**

#### Least Absolute Shrinkage and Selection Operator (Lasso)

The optimization problem of Lasso can be written as

$$\min_{\bm{w}} \frac{1}{2} \| \bm{y} - \bm{X}^{\top} \bm{w} \|_2^2 + \lambda \| \bm{w} \|_1,$$

where  $\|\boldsymbol{w}\|_1 = \sum_{k=1}^d |w_k|$  is an  $\ell_1$  norm.

Lasso is a convex method: The first term is a convex function w.r.t. w.  $\ell_1$  norm (all norm) is convex:

$$\|\alpha w + (1 - \alpha)v\|_1 \le \|\alpha w\|_1 + \|(1 - \alpha)v\|_1$$
  
=  $\alpha \|w\|_1 + (1 - \alpha)\|v\|_1$ 

where  $0 \le \alpha \le 1$ . The sum of two convex functions is convex.

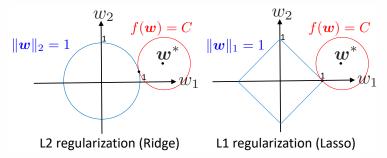
## **Embedded Method (Lasso)**

The  $\ell_1$  regularization is equivalent to  $\ell_1$  norm constraint:

$$\min_{\boldsymbol{w}} \quad f(\boldsymbol{w}) + \lambda \|\boldsymbol{w}\|_1 \longrightarrow \min_{\boldsymbol{w}} \quad f(\boldsymbol{w}), \quad \text{s.t.} \quad \|\boldsymbol{w}\|_1 \leq \eta.$$

There exists the same solution of the  $\ell_1$  norm constraint with an arbitrary  $\lambda$ .

Using the  $\ell_1$  regularizer, we can make w sparse.



## When Lasso helpful?

Let us think about a least-squared regression problems:

$$\min_{oldsymbol{w} \in \mathbb{R}^d} \ \|oldsymbol{y} - oldsymbol{X}^ op oldsymbol{w}\|_2^2.$$

Take the objective function with respect to  $oldsymbol{w}$  and set it to zero:

$$\frac{\partial}{\partial \boldsymbol{w}} \|\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}\|_{2}^{2} = -2\boldsymbol{X}(\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}) = \boldsymbol{0}$$

Use Eq. (84) of [1]. The solution is given as

$$\widehat{\boldsymbol{w}} = (\boldsymbol{X} \boldsymbol{X}^{\top})^{-1} \boldsymbol{X} \boldsymbol{y}.$$

If the rank of  $\boldsymbol{X}$  is d, the rank of  $\boldsymbol{X}\boldsymbol{X}^{\top}$  is also d and it is invertible.

What happens if the rank of X is less than d?

## Lasso with ADMM (1/8)

Lasso has no closed form solution. Thus, we need to iteratively optimize the problem.

Here, we introduce the Alternating Direction Method of Multipliers (ADMM) [5].

We can rewrite the Lasso optimization problem as

$$\min_{\boldsymbol{w},\boldsymbol{z}} \ \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}\|_{2}^{2} + \lambda \|\boldsymbol{z}\|_{1} + \frac{\rho}{2} \|\boldsymbol{w} - \boldsymbol{z}\|_{2}^{2}$$
s.t.  $\boldsymbol{w} = \boldsymbol{z}$ 

The key idea here is to split the main objective and the non-differentiable regularization term. Since the last term  $\frac{\rho}{2} \| \boldsymbol{w} - \boldsymbol{z} \|_2^2$  is zero if the constraint is satisfied, this problem is equivalent to the original Lasso problem.

27 / 36

## Lasso with ADMM (2/8)

Let us denote the Lagrange multipliers as  $\gamma \in \mathbb{R}^d$ , we can write a Lagrangian function (called Augmented Lagrangian function) as follows:

$$J(w, z, \gamma) = \frac{1}{2} \|y - X^{\top} w\|_{2}^{2} + \gamma^{\top} (w - z)$$
  
  $+ \lambda \|z\|_{1} + \frac{\rho}{2} \|w - z\|_{2}^{2},$ 

where  $\rho > 0$  is a tuning parameter.

## Lasso with ADMM (3/8)

In ADMM, we consider the following optimization problem:

$$egin{aligned} \min_{oldsymbol{w},oldsymbol{z}} \max_{oldsymbol{\gamma}} & J(oldsymbol{w},oldsymbol{z},oldsymbol{\gamma}) = rac{1}{2}\|oldsymbol{y} - oldsymbol{X}^ op oldsymbol{w}\|_2^2 + oldsymbol{\gamma}^ op (oldsymbol{w} - oldsymbol{z}) \ & + \lambda \|oldsymbol{z}\|_1 + rac{
ho}{2}\|oldsymbol{w} - oldsymbol{z}\|_2^2, \end{aligned}$$

Since we have the relationship,

$$\max_{\boldsymbol{\gamma}} J(\boldsymbol{w}, \boldsymbol{z}, \boldsymbol{\gamma}) = \begin{cases} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}\|_{2}^{2} + \lambda \|\boldsymbol{z}\|_{1} & (\boldsymbol{w} = \boldsymbol{z}) \\ \infty & (\text{Otherwise}) \end{cases}$$

The optimization problem is equivalent to the original Lasso problem.

## Lasso with ADMM (4/8)

Minimizing  $J(\boldsymbol{w}, \boldsymbol{z}, \boldsymbol{\gamma})$  w.r.t.  $\boldsymbol{w}$ . If we fix  $\boldsymbol{z}$  and  $\boldsymbol{\gamma}$  as  $\boldsymbol{z}^{(t)}$  and  $\boldsymbol{\gamma}^{(t)}$ ,  $J(\boldsymbol{w}, \boldsymbol{z}^{(t)}, \boldsymbol{\gamma}^{(t)})$  is convex w.r.t.  $\boldsymbol{w}$ . That is,

$$\frac{\partial J(\boldsymbol{w},\boldsymbol{z},\boldsymbol{\gamma})}{\partial \boldsymbol{w}} = -\boldsymbol{X}(\boldsymbol{y} - \boldsymbol{X}^{\top}\boldsymbol{w}) + \boldsymbol{\gamma} + \rho(\boldsymbol{w} - \boldsymbol{z}) = 0.$$

Here, we can use the following equation (see [1] Eq. (84)):

$$\frac{\partial \|\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}\|_{2}^{2}}{\partial \boldsymbol{w}} = -2\boldsymbol{X}(\boldsymbol{y} - \boldsymbol{X}^{\top} \boldsymbol{w}).$$

Solving it for w:

$$(XX^{\top} + \rho I)w = Xy - \gamma^{(t)} + \rho z^{(t)}$$
  
$$w^{(t+1)} = (XX^{\top} + \rho I)^{-1}(Xy - \gamma^{(t)} + \rho z^{(t)}).$$

## Lasso with ADMM (5/8)

Minimizing  $J(w,z,\gamma)$  w.r.t. z. If we fix w and  $\gamma$  as  $w^{(t)}$  and  $\gamma^{(t)}$ ,  $J(w^{(t)},z,\gamma^{(t)})$  is convex w.r.t. z.

$$J(\boldsymbol{w}^{(t)}, \boldsymbol{z}, \boldsymbol{\gamma}^{(t)}) = \frac{\rho}{2} \|\boldsymbol{z} - \boldsymbol{w}^{(t)}\|_2^2 + \lambda \|\boldsymbol{z}\|_1 - \boldsymbol{\gamma}^\top \boldsymbol{z} + \text{Const.}$$

 $||z||_1$  is not differentiable at 0. However, we can analytically solve the problem! Moreover, since there is no interaction in the elements of z, we can solve it for each element.

$$J(\boldsymbol{w}^{(t)}, (z_1, \dots, z_\ell, \dots, z_d), \boldsymbol{\gamma}^{(t)}) = \frac{\rho}{2} (z_\ell - w_\ell^{(t)})^2 + \lambda |z_\ell| - \gamma_\ell z_\ell + \mathsf{Const.}$$

## Lasso with ADMM (6/8)

#### Case1:

$$z_{\ell} > 0, \rho(z_{\ell} - w_{\ell}^{(t)}) + \lambda - \gamma_{\ell} = 0 \longrightarrow z_{\ell} = w_{\ell}^{(t)} + \frac{1}{\rho}(\gamma_{\ell} - \lambda)$$

That is, 
$$z_\ell > 0$$
 if  $w_\ell^{(t)} + \frac{1}{\rho} \gamma_\ell > \frac{\lambda}{\rho}$ 

#### Case2:

$$z_\ell < 0, 
ho(z_\ell - w_\ell^{(t)}) - \lambda - \gamma_\ell = 0 \longrightarrow z_\ell = w_\ell^{(t)} + rac{1}{
ho}(\gamma_\ell + \lambda)$$

That is, 
$$z_{\ell} < 0$$
 if  $w_{\ell}^{(t)} + \frac{1}{\rho} \gamma_{\ell} < -\frac{\lambda}{\rho}$ 

Case3: 
$$z_{\ell} = 0$$
,  $0 \in \rho(z_{\ell} - w_{\ell}^{(t)}) + \lambda[-1 \ 1] - \gamma_{\ell} \longrightarrow w_{\ell} + \frac{1}{\rho}\gamma_{\ell} \in [-\frac{\lambda}{\rho}, \frac{\lambda}{\rho}], (z_{\ell} = 0).$ 

## Lasso with ADMM (7/8)

Let us introduce the Soft-Thresholding function:

$$S_{\lambda}(x) = \begin{cases} x - \lambda & (x > \lambda) \\ 0 & (x \in [-\lambda, \lambda]) \\ x + \lambda & (x < -\lambda) \end{cases}$$
$$= \operatorname{sign}(x) \max(0, |x| - \lambda)$$

Therefore, the update of  $z_\ell$  can be simply written by the soft-thresholding function as

$$\widehat{z}_{\ell}^{(t+1)} = S_{\frac{\lambda}{\rho}} \left( w_{\ell}^{(t)} + \frac{1}{\rho} \gamma_{\ell} \right).$$

## Lasso with ADMM (8/8)

Maximizing  $J(w,z,\gamma)$  w.r.t.  $\gamma$ . That is the optimization problem can be written as

$$\max_{oldsymbol{\gamma}} J(oldsymbol{w}, oldsymbol{z}, oldsymbol{\gamma}) = oldsymbol{\gamma}^{ op}(oldsymbol{w} - oldsymbol{z}).$$

To optimize this problem, since we cannot get the analytical solution, we use the gradient ascent algorithm:

$$\gamma^{(t+1)} = \gamma^{(t)} + \rho(w^{(t)} - z^{(t)}).$$

Thus, the ADMM algorithm for Lasso can be summarized as

$$egin{aligned} oldsymbol{w}^{(t+1)} &= (oldsymbol{X}oldsymbol{X}^{ op} + 
ho oldsymbol{I})^{-1} (oldsymbol{X}oldsymbol{y} - oldsymbol{\gamma}^{(t)} + 
ho oldsymbol{Z}^{(t)}) \ oldsymbol{z}^{(t+1)}_{\ell} &= S_{rac{\lambda}{
ho}} (oldsymbol{w}^{(t+1)} + rac{1}{
ho}oldsymbol{\gamma}) \ oldsymbol{\gamma}^{(t+1)} &= oldsymbol{\gamma}^{(t+1)} + 
ho (oldsymbol{w}^{(t+1)} - oldsymbol{z}^{(t+1)}). \end{aligned}$$

#### **Elastic-Net**

For Lasso, the number of non-zero features should be smaller than n. How to select r > n variables?

Ans: Use the elastic net regularization [6]:

$$\min_{\bm{w}} \ \|\bm{y} - \bm{X}^{\top} \bm{w}\|_2^2 + \lambda (\alpha \|\bm{w}\|_1 + (1 - \alpha) \|\bm{w}\|_2^2),$$

where  $0 \le \alpha \le 1$  and  $\lambda > 0$  is a regularization parameter.

 $\|\boldsymbol{w}\|_2^2$  is differentiable; we can similarly solve it with ADMM.

$$\begin{aligned} \boldsymbol{w}^{(t+1)} &= (\boldsymbol{X}\boldsymbol{X}^{\top} + 2\lambda(1-\alpha)\boldsymbol{I} + \rho\boldsymbol{I})^{-1}(\boldsymbol{X}\boldsymbol{y} - \boldsymbol{\gamma}^{(t)} + \rho\boldsymbol{z}^{(t)}) \\ \boldsymbol{z}_{\ell}^{(t+1)} &= S_{\frac{\lambda\alpha}{\rho}}(\boldsymbol{w}^{(t+1)} + \frac{1}{\rho}\boldsymbol{\gamma}) \\ \boldsymbol{\gamma}^{(t+1)} &= \boldsymbol{\gamma}^{(t+1)} + \rho(\boldsymbol{w}^{(t+1)} - \boldsymbol{z}^{(t+1)}). \end{aligned}$$

Thanks to the  $\ell_2$  regularization, w tends to be dense.

## Summary

- Feature selection: Wrapper method, Filter method, and Embedded method
- Wrapper method (Selecting features that maximize prediction accuracy. Computationally expensive.)
- Filter method (Use mutual information to select features, e.g., MR, mRMR, etc.)
- Embedded method (Selecting features during training. e.g., Lasso)
- Alternating Direction Method of Multipliers (ADMM).

- [1] Kaare Brandt Petersen, Michael Syskind Pedersen, et al. The matrix cookbook. *Technical University of Denmark*, 7:15, 2008.
- [2] H. Peng, F. Long, and C. Ding. Feature selection based on mutual information: Criteria of max-dependency, max-relevance, and min-redundancy. *IEEE Transactions on Pattern Analysis and Machine Intelligence*, 27:1226–1237, 2005.
- [3] A. Gretton, O. Bousquet, Alex. Smola, and B. Schölkopf. Measuring statistical dependence with Hilbert-Schmidt norms. In ALT, 2005.
- [4] C. Cortes, M. Mohri, and A. Rostamizadeh. Algorithms for learning kernels based on centered alignment. *JMLR*, 13:795–828, 2012.
- [5] Stephen Boyd, Neal Parikh, Eric Chu, Borja Peleato,

- Jonathan Eckstein, et al. Distributed optimization and statistical learning via the alternating direction method of multipliers. Foundations and Trends $\mathbb{R}$  in Machine learning, 3(1):1–122, 2011.
- [6] Hui Zou and Trevor Hastie. Regularization and variable selection via the elastic net. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 67(2):301–320, 2005.